

Meta-Regression

Subgroup Analysis

This vignette introduces the regression functionality of the `dtametaTMB` package for meta-analysis of diagnostic test accuracy (DTA) studies.

Validation

The subgroup HSROC and Reitsma implementations reproduce the Cochrane Handbook RF and Anti-CCP examples closely, including the baseline parameters (HSROC: accuracy, threshold, shape; Reitsma: logit sensitivity and specificity), between-study variance estimates, and subgroup effects on accuracy and threshold. The Schuetz CT/MRI analyses yielded results broadly consistent with the published subgroup parameter estimates and variance components. Likelihood-ratio tests led to identical substantive conclusions, although the LR statistics differed somewhat from the published values.

Dummy-coding vs. cell-means parameterization

For subgroup analyses, we distinguish between two equivalent parameterizations of the same model. In the reference-group parameterization, one subgroup serves as the baseline and the remaining subgroup effects are expressed as deviations from this reference. In the group-specific (cell-means) parameterization, each subgroup has its own parameter directly. The former is convenient for testing subgroup differences, whereas the latter is convenient for reporting subgroup-specific estimates and plotting subgroup-specific HSROC curves. The Reitsma subgroup model is fitted twice with both parameterizations. The subgroup HSROC model is estimated using a reference-group parameterization, and subgroup-specific parameters are recovered by evaluating the fitted linear predictors at subgroup-specific design points via Z_{pred} .

How do we include covariates in the Reitsma model?

For sensitivity in study i , we have the number of diseased individuals testing positive: $y_{Ai} \sim \mathcal{B}(n_{Ai}, \pi_{Ai})$.

Similarly for specificity, we have the number of non-diseased individuals testing negative: $y_{Bi} \sim \mathcal{B}(n_{Bi}, \pi_{Bi})$.

Now we introduce a p -dimensional design-vector z_i including study level covariates. Consequently, at the study level, we have

$$\begin{pmatrix} z_i^\top \mu_{Ai} \\ z_i^\top \mu_{Bi} \end{pmatrix} \sim \mathcal{N} \left(\begin{pmatrix} z_i^\top \mu_A \\ z_i^\top \mu_B \end{pmatrix}, \Sigma \right), \quad \text{with} \quad \Sigma = \begin{pmatrix} \sigma_A^2 & \sigma_{AB} \\ \sigma_{AB} & \sigma_B^2 \end{pmatrix}.$$

Let's assume that z_i includes a single binary covariate representing two subgroups. Then using dummy coding with subgroup 1 being the reference, we have

$$z_i^\top \mu_{Ai} = \begin{cases} \mu_{Ai} & \text{for subgroup 1,} \\ \mu_{Ai} + \nu_{A2} & \text{for subgroup 2,} \end{cases} \quad z_i^\top \mu_{Bi} = \begin{cases} \mu_{Bi} & \text{for subgroup 1,} \\ \mu_{Bi} + \nu_{B2} & \text{for subgroup 2,} \end{cases}$$

$$z_i^\top \mu_A = \begin{cases} \mu_A & \text{for subgroup 1,} \\ \mu_A + \nu_{A2} & \text{for subgroup 2,} \end{cases} \quad z_i^\top \mu_B = \begin{cases} \mu_B & \text{for subgroup 1,} \\ \mu_B + \nu_{B2} & \text{for subgroup 2.} \end{cases}$$

```
data("anticcp")
reitsmasub <- fitReitsmaSubgroup(data=anticcp,
                                TP=TP,
                                FP=FP,
                                FN=FN,
                                TN=TN,
                                study=study,
                                subgroup=generation)

reitsmasub
#>
#> Reitsma Subgroup Model
#> -----
#>
#> Number of studies      : 37
#> Number of subgroups   : 2
#> Optimizer              : Converged
#> Hessian                : Positive definite
#> Max |grad|             : 0.0004858134
#> -2 log likelihood      : 533.37 ( df = 7 )
```

```

#> AIC : 547.37
#> BIC : 558.646
#>
#> Use summary() for parameter estimates.
summary(reitsmasub)
#> $estimates
#>
#> Estimate Std_Error
#> mu_A.CCP1 -0.09653883 0.22032058
#> mu_B.CCP1 3.44671920 0.29824368
#> mu_A.CCP2 0.86603855 0.12087681
#> mu_B.CCP2 3.01649066 0.16223753
#> sigma2_A.sens 0.35982866 0.10217649
#> sigma2_B.spec 0.53990927 0.18015721
#> sigma_AB -0.19684497 0.09835341
#> nu_A.CCP2 0.96257151 0.25134200
#> nu_B.CCP2 -0.43020982 0.33771185
#>
#> $sensspec
#>
#> type Estimate conflevel CI_Lower CI_Upper
#> mu_A.CCP1 sens 0.4758840 0.95 0.3708997 0.5830439
#> mu_B.CCP1 spec 0.9691331 0.95 0.9459445 0.9825578
#> mu_A.CCP2 sens 0.7039207 0.95 0.6522909 0.7508130
#> mu_B.CCP2 spec 0.9533136 0.95 0.9369387 0.9655926
#>
#> $RutterGatsonis_recovered
#>
#> Lambda Theta beta sigma2_alpha sigma2_theta
#> CCP1 3.007377 -1.6105343 0.2028866 0.4878424 0.3188056
#> CCP2 3.684000 -0.8834971 0.2028866 0.4878424 0.3188056
#>
#> $subgroups
#> [1] "CCP1" "CCP2"

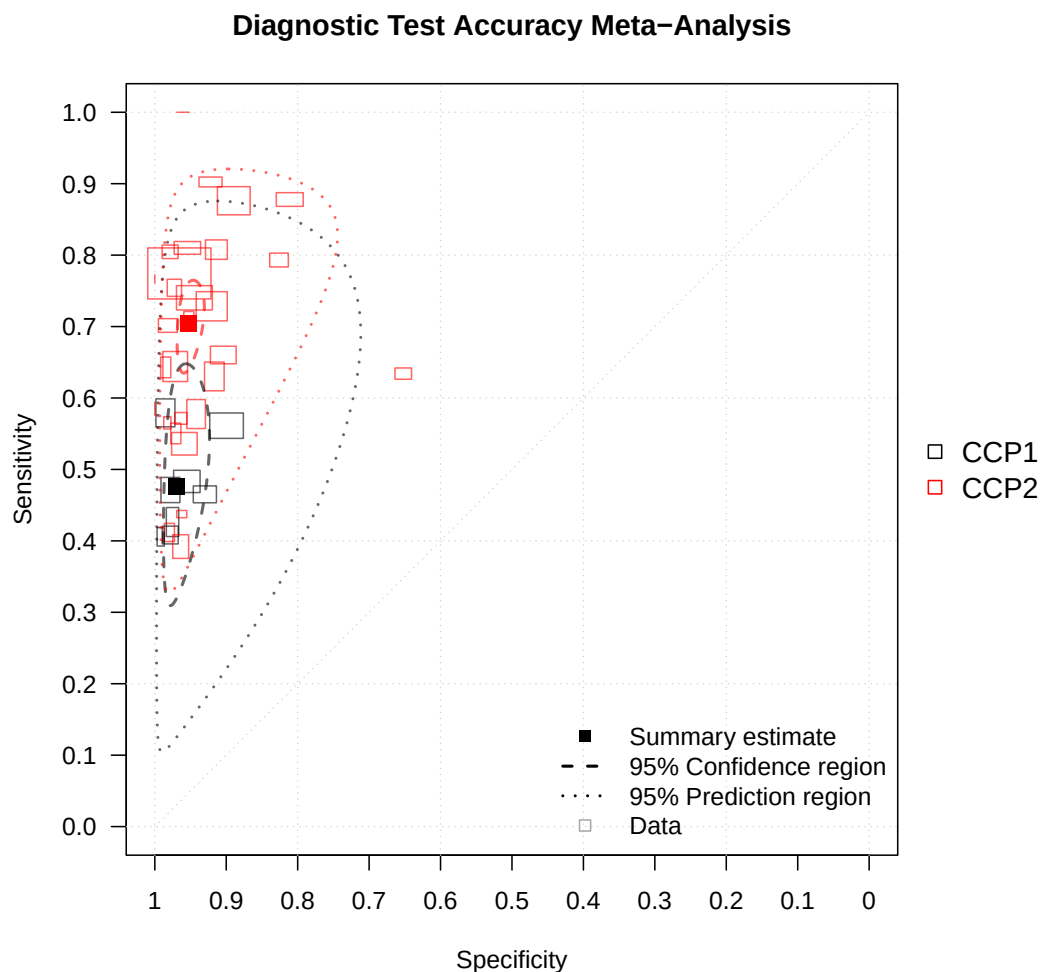
```

How do I get a summary plot of the Reitsma model?

```

plot(reitsmasub,
     scale=0.01,
     nudge_legend=-0.2,
     size="se",
     col=c("black","red"))

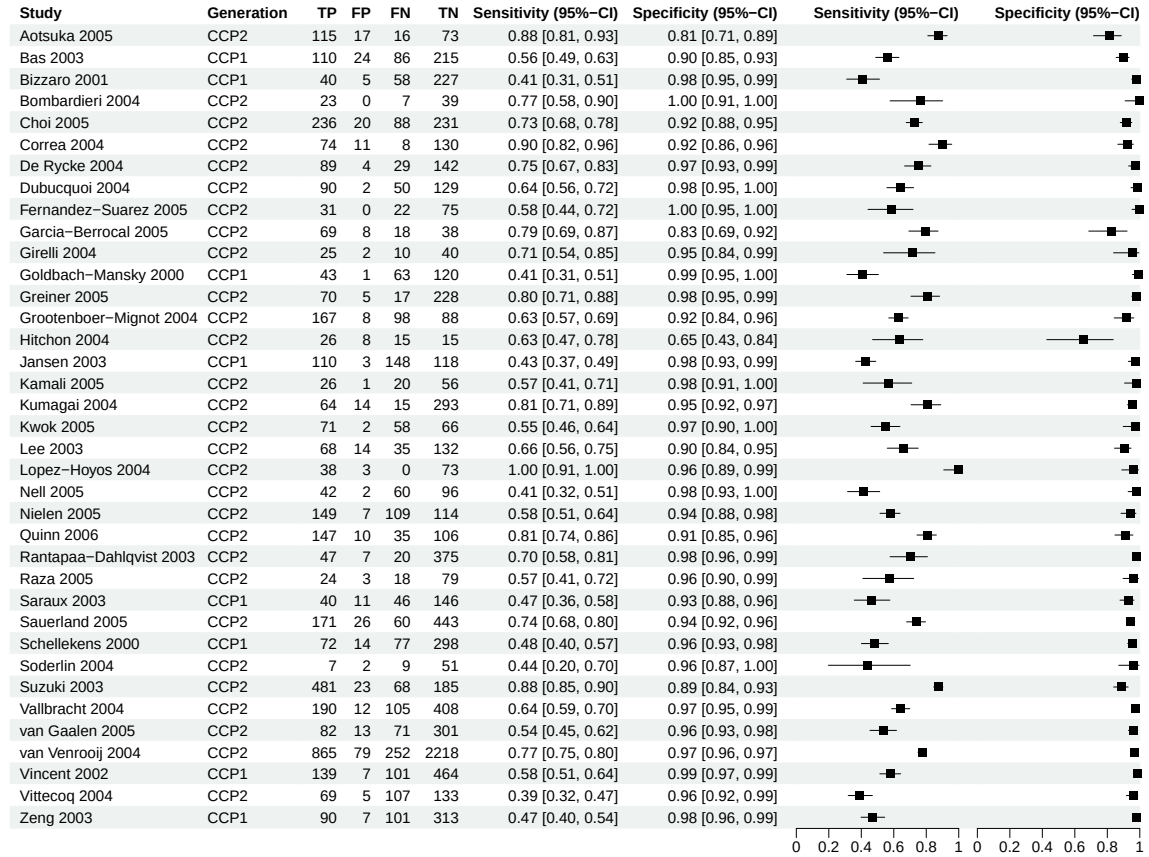
```



How do I get a coupled forest plot?

Note: Rendering forest plots may take longer in an interactive R session due to the underlying grid graphics. Performance is typically faster when knitting the vignette (e.g. via RMarkdown or Quarto).

```
forest(reitsmasub, subgroup_label="Generation")
```



How do I constrain parameters in the Reitsma model?

In sparse data one may wish to fix parameters of the random effects (variance-covariance matrix) at zero. This can be done via the `constrain` argument.

For example,

```
constrainA <- fitReitsmaSubgroup(data=anticcp,
                                TP=TP,
                                FP=FP,
                                FN=FN,
                                TN=TN,
                                study=study,
```

```

                                subgroup=generation,
                                constrain="sigma2_A")
summary(constrainA)$estimates
#>               Estimate Std_Error
#> mu_A.CCP1      -5.439405e-02 0.05498530
#> mu_B.CCP1       3.446625e+00 0.29875482
#> mu_A.CCP2       9.179794e-01 0.03139172
#> mu_B.CCP2       2.996859e+00 0.16216233
#> sigma2_A.sens   4.930381e-32 0.00000000
#> sigma2_B.spec   5.396331e-01 0.17997640
#> sigma_AB        0.000000e+00 0.00000000
#> nu_A.CCP2       9.723735e-01 0.06331527
#> nu_B.CCP2      -4.497661e-01 0.33795382

```

fixes the logit sensitivity variance to zero. This also implies that the random effects covariance is zero. Note that you can also set `constrain` to "sigma_AB", "sigma2_B", or "all".

Constraining fixed effects is controlled by the `sensspec_constrain` argument. For example,

```

constrainsens <- fitReitsmaSubgroup(data=anticcp,
                                   TP=TP,
                                   FP=FP,
                                   FN=FN,
                                   TN=TN,
                                   study=study,
                                   subgroup=generation,
                                   sensspec_constrain="sens")
summary(constrainsens)$estimates
#>               Estimate Std_Error
#> mu_A.CCP1      0.65330520 0.1274076
#> mu_B.CCP1      3.08122030 0.3256178
#> mu_A.CCP2      0.65330520 0.0000000
#> mu_B.CCP2      3.11800628 0.1745128
#> sigma2_A.sens   0.54192926 0.1462735
#> sigma2_B.spec   0.57765267 0.1996838
#> sigma_AB       -0.27715832 0.1398143
#> nu_A.CCP2       0.00000000 0.0000000
#> nu_B.CCP2       0.03678907 0.3857733

```

assumes equal (logit) sensitivities in all subgroups. Note that you can also set `subgroup_constrain` to "spec" or `c("sens", "spec")`.

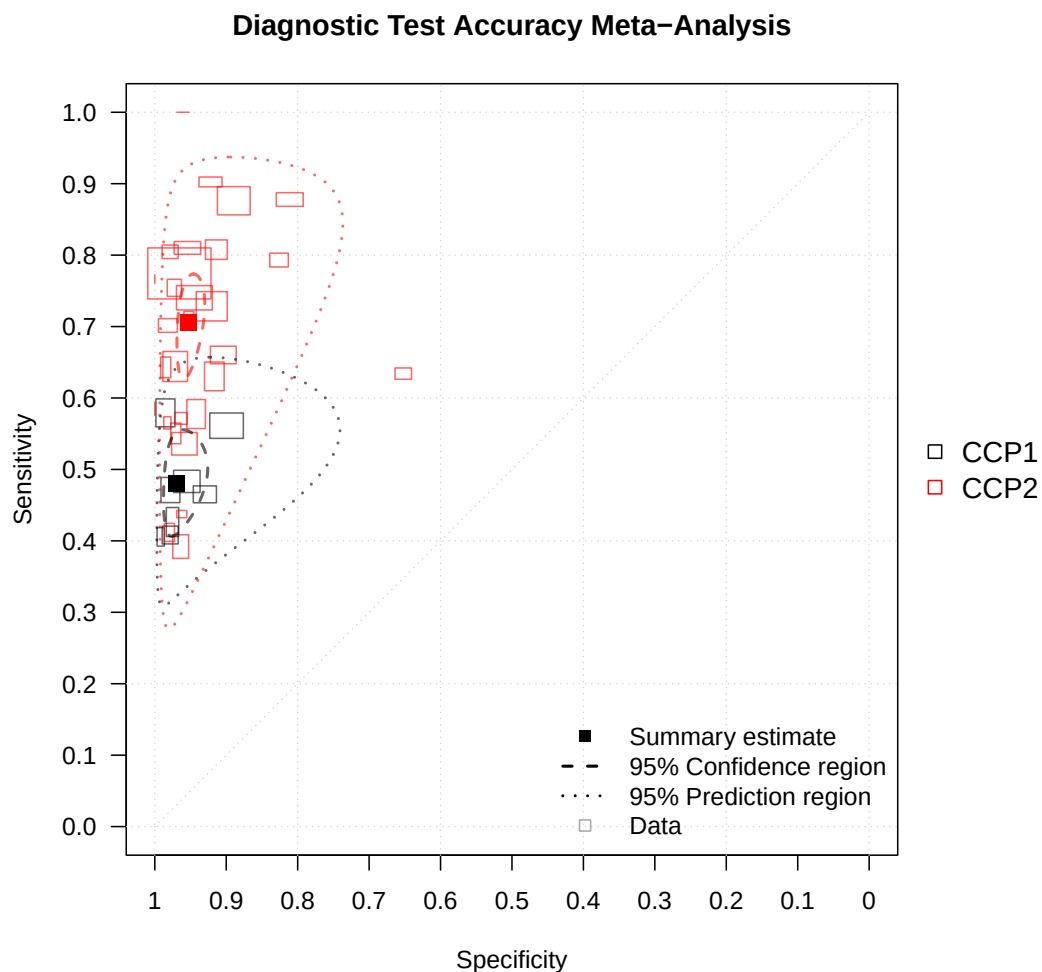
How can I compare constrainsens with the full model?

We can perform a likelihood ratio test via `anova`, essentially testing whether there exist subgroup differences in sensitivity.

```
anova(constrainsens,reitsmasub)
#>           Df  logLik Df.diff  Chisq Pr(>Chisq)
#> Model 1   6 -272.77
#> Model 2   7 -266.69       1 12.181  0.0004827 ***
#> ---
#> Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

How do I allow for different subgroup-specific random-effects (co-)variances in the Reitsma model?

```
heteroskedastic <- fitReitsmaSubgroup(data=anticcp,
                                     TP=TP,
                                     FP=FP,
                                     FN=FN,
                                     TN=TN,
                                     study=study,
                                     subgroup=generation,
                                     variances="unequal")
plot(heteroskedastic,
     scale=0.01,
     nudge_legend=-0.2,
     size="se",
     col=c("black","red"))
```



```
anova(reitsmasub,heteroskedastic)
#>      Df logLik Df.diff  Chisq Pr(>Chisq)
#> Model 1  7 -266.69
#> Model 2 10 -262.42      3 8.5306   0.03623 *
#> ---
#> Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

How do we include covariates in the Rutter and Gatsonis model?

The number of diseased individuals from study i who test positive is denoted by $y_{i1} \sim \mathcal{B}(n_{i1}, \pi_{i1})$.

Similarly, the number of non-diseased individuals who test positive is $y_{i2} \sim \mathcal{B}(n_{i2}, \pi_{i2})$.

Now we introduce a p -dimensional design-vector z_i including study level covariates. Consequently, at the study level, we have

$$\text{logit}(\pi_{ij}) = (z_i^\top \theta_i + z_i^\top \alpha_i x_{ij}) \exp(-z_i^\top \beta x_{ij}),$$

$$z_i^\top \alpha_i \sim \mathcal{N}(z_i^\top \Lambda, \sigma_\alpha^2), \quad z_i^\top \theta_i \sim \mathcal{N}(z_i^\top \Theta, \sigma_\theta^2),$$

where

$$x_{ij} = \begin{cases} -0.5 & \text{for non-diseased individuals,} \\ 0.5 & \text{for diseased individuals.} \end{cases}$$

Let's assume that z_i includes a single categorical covariate representing three subgroups. Then using dummy coding with subgroup 1 being the reference, we have

$$\begin{aligned} z_i^\top \theta_i &= \begin{cases} \theta_i & \text{for subgroup 1,} \\ \theta_i + \gamma_2 & \text{for subgroup 2,} \\ \theta_i + \gamma_3 & \text{for subgroup 3,} \end{cases} & z_i^\top \alpha_i &= \begin{cases} \alpha_i & \text{for subgroup 1,} \\ \alpha_i + \xi_2 & \text{for subgroup 2,} \\ \alpha_i + \xi_3 & \text{for subgroup 3,} \end{cases} \\ z_i^\top \beta &= \begin{cases} \beta & \text{for subgroup 1,} \\ \beta + \delta_2 & \text{for subgroup 2,} \\ \beta + \delta_3 & \text{for subgroup 3,} \end{cases} \\ z_i^\top \Lambda &= \begin{cases} \Lambda & \text{for subgroup 1,} \\ \Lambda + \xi_2 & \text{for subgroup 2,} \\ \Lambda + \xi_3 & \text{for subgroup 3,} \end{cases} & z_i^\top \Theta_i &= \begin{cases} \Theta & \text{for subgroup 1,} \\ \Theta + \gamma_2 & \text{for subgroup 2,} \\ \Theta + \gamma_3 & \text{for subgroup 3.} \end{cases} \end{aligned}$$

Note: For prediction `fitRutterGatsonisSubgroup()` uses the prediction matrix

$$Z_{\text{pred}} = \begin{pmatrix} 1 & 0 & 0 \\ 1 & 1 & 0 \\ 1 & 0 & 1 \end{pmatrix}$$

for the case above and therefore recovers the threshold, accuracy, and shape parameters for each subgroup as if it were the reference group.

How do I perform a Rutter and Gatsonis meta-regression with a single categorical study-level covariate?

```

data("RF")
RF2      <- RF[RF$method %in% c("LA","ELISA","Nephelometry"),]
RF2$method <- factor(RF2$method,levels=c("LA","ELISA","Nephelometry"))
ruttergatsonissub <- fitRutterGatsonisSubgroup(data=RF2,
                                              TP=TP,
                                              FP=FP,
                                              FN=FN,
                                              TN=TN,
                                              study=study,
                                              subgroup=method,
                                              constrain="shape") # assumes equal
                                                                # shapes in subgroups

```

```

ruttergatsonissub
#>
#> Rutter & Gatsonis Subgroup Model
#> -----
#>
#> Number of studies      : 47
#> Number of subgroups   : 3
#> Optimizer              : Converged
#> Hessian                : Positive definite
#> Max |grad|             : 0.0001105815
#> -2 log likelihood      : 753.722 ( df = 9 )
#> AIC                   : 771.722
#> BIC                   : 788.374
#>
#> Use summary() for parameter estimates.
summary(ruttergatsonissub)
#> $estimates
#>
#>               Estimate Std. Error
#> Lambda_LA      2.45630477 0.32357688
#> xi_ELISA       0.24853315 0.43954107
#> xi_Nephelometry 0.33142788 0.44260056
#> Theta_LA      -0.55005213 0.21334028
#> gamma_ELISA    -0.19625486 0.26096194
#> gamma_Nephelometry 0.49585945 0.26223043
#> beta_LA        0.19768573 0.16983959
#> delta_ELISA     0.00000000 0.00000000
#> delta_Nephelometry 0.00000000 0.00000000
#> sigma2_alpha    1.27791458 0.30852474

```

```

#> sigma2_theta      0.47684025 0.11344515
#> Lambda_LA         2.45630477 0.32357688
#> Lambda_ELISA      2.70483792 0.32695410
#> Lambda_Nephelometry 2.78773265 0.30577909
#> Theta_LA          -0.55005213 0.21334028
#> Theta_ELISA       -0.74630699 0.20991299
#> Theta_Nephelometry -0.05419268 0.21213124
#> beta_LA           0.19768573 0.16983959
#> beta_ELISA         0.19768573 0.16983959
#> beta_Nephelometry 0.19768573 0.16983959
#> logitsens         0.82616623 0.28629916
#> logitsens         1.05130869 0.28606874
#> logitsens         1.12640178 0.27689464
#> sens              0.69554369 0.06062747
#> sens              0.74102612 0.05489842
#> sens              0.75517425 0.05119397
#>
#> $sensspec
#>      subgroup      spec conflevel logitsens Std_Error CI_Lower CI_Upper
#> 1          LA 0.8461538      0.95 0.8261662 0.2862992 0.2650302 1.387302
#> 2          ELISA 0.8461538      0.95 1.0513087 0.2860687 0.4906243 1.611993
#> 3 Nephelometry 0.8461538      0.95 1.1264018 0.2768946 0.5836983 1.669105
#>      Sens SensCI_Lower SensCI_Upper
#> 1 0.6955437      0.5658724      0.8001612
#> 2 0.7410261      0.6202535      0.8336879
#> 3 0.7551742      0.6419179      0.8414565
#>
#> $Reitsma_recovered
#>      mu_A.sens mu_B.spec sigma2_A.sens sigma2_B.spec sigma_AB
#> LA          0.6142809 1.962947      0.6534814      0.9703778 -0.1573616
#> ELISA        0.5490678 2.316769      0.6534814      0.9703778 -0.1573616
#> Nephelometry 1.2135903 1.598502      0.6534814      0.9703778 -0.1573616
#>
#> $subgroups
#> [1] "LA"          "ELISA"         "Nephelometry"

```

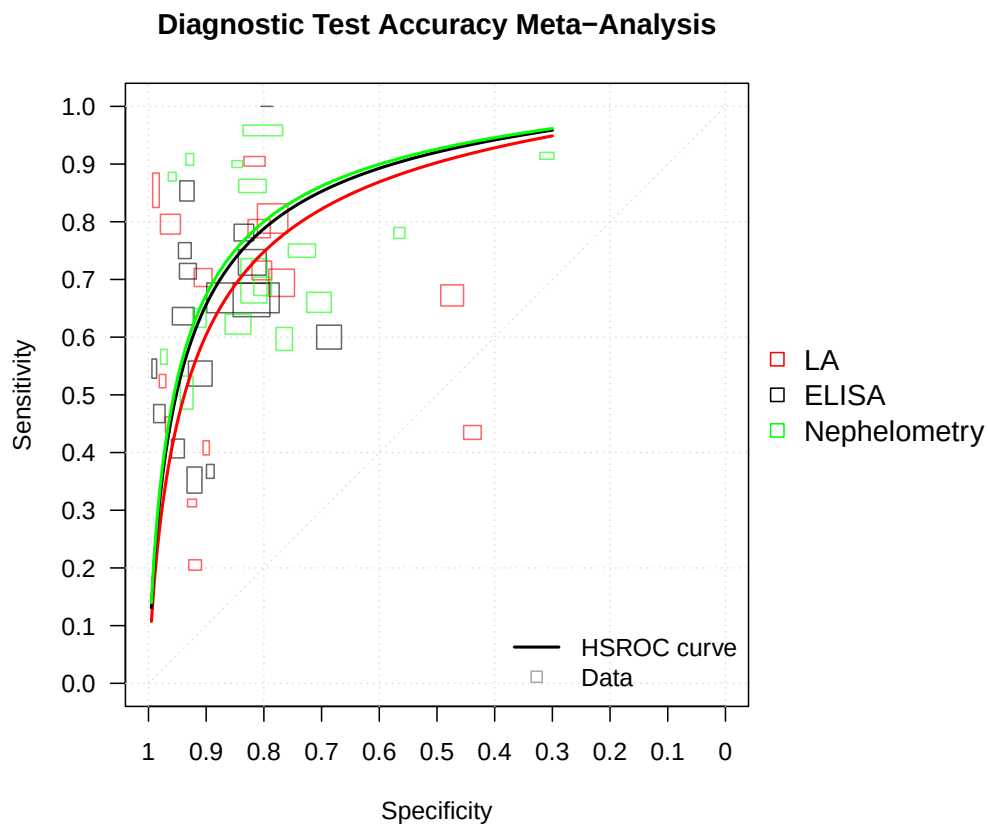
How do I get a summary plot of the Rutter and Gatsonis subgroup model?

```

plot(ruttergatsonissub,
     specrange=c(0.3,0.995),

```

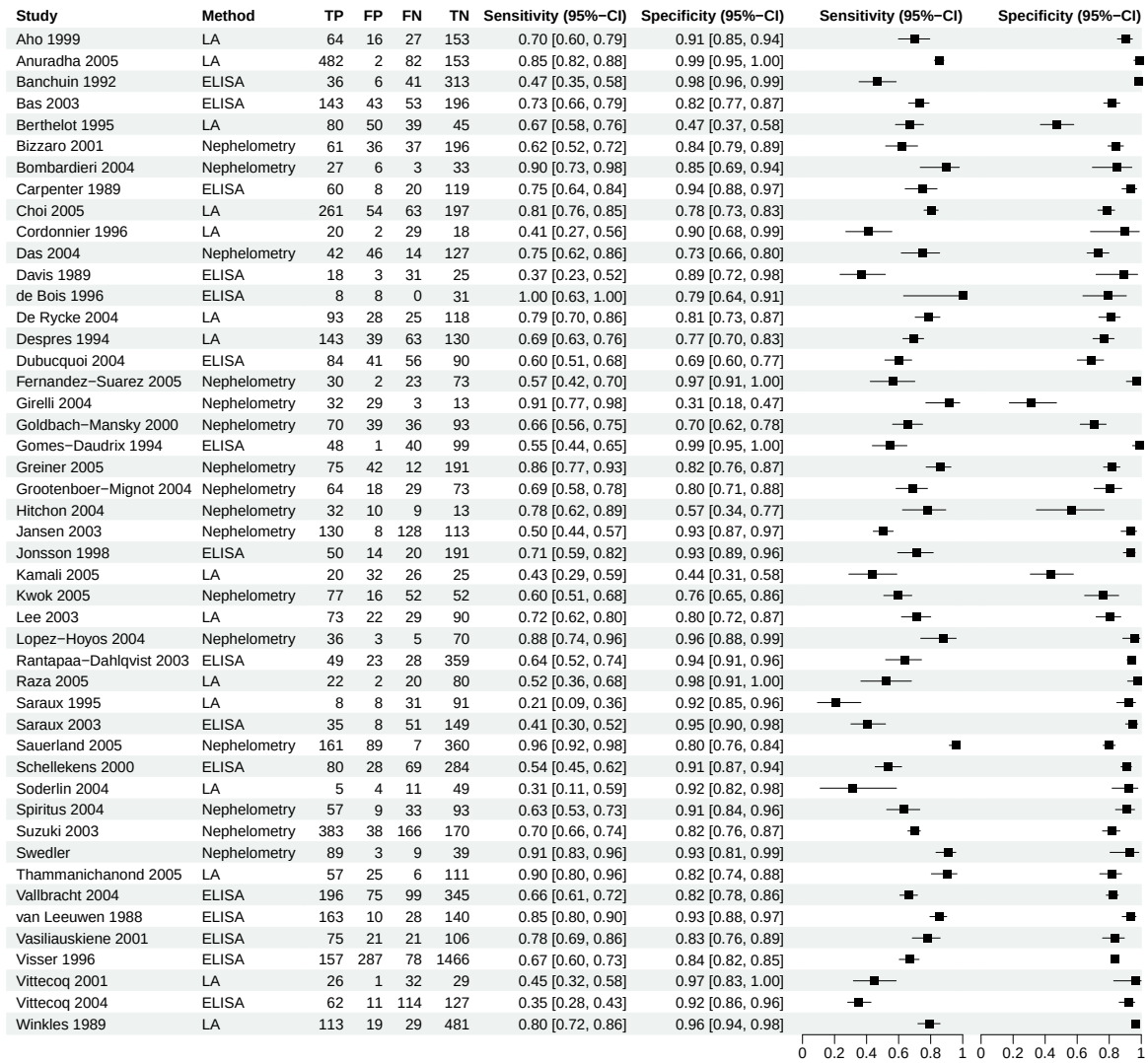
```
size="se",
col=c("red","black","green"),
scale=0.015)
```



How do I get a coupled forest plot?

Note: Rendering forest plots may take longer in an interactive R session due to the underlying grid graphics. Performance is typically faster when knitting the vignette (e.g. via RMarkdown or Quarto).

```
forest(ruttergatsonissub, subgroup_label = "Method")
```



How do I constrain parameters in the Rutter and Gatsonis subgroup model?

In the Rutter and Gatsonis model, all parameter constraints are controlled by `constrain`, i.e.,

- `constrain="sigma2_alpha"` sets σ_{α}^2 to zero,
- `constrain="sigma2_theta"` sets σ_{θ}^2 to zero,

- `constrain="accuracy"` assumes equal accuracy parameters across subgroups, i.e. $\xi_2 = \xi_3 = \dots = 0$,
- `constrain="threshold"` assumes equal threshold parameters across subgroups, i.e. $\gamma_2 = \gamma_3 = \dots = 0$,
- `constrain="shape"` assumes equal shape parameters across subgroups, i.e. $\delta_2 = \delta_3 = \dots = 0$,
- `constrain="shape_zero"` fixes all shape parameters at zero.

Constraints can also be combined, for example `constrain=c("shape", "sigma2_theta")`.

How do I use the general Rutter and Gatsonis regression function?

This method is for advanced users who feel comfortable specifying their own design and prediction matrices. For study-level covariates, the design matrix Z needs two identical consecutive rows per study, one for the diseased and one for the non-diseased. Of note, there are neither `plot()` nor `forest()` methods for `fitRutterGatsonisReg()`.

Let's reproduce the subgroup-analysis from before.

```
# Specify design matrix Z
Z <- model.matrix(~method, data=RF2)
# For study level-covariates, we need to two identical consecutive
# rows per study (diseased and non-diseased).
Z2 <- Z[rep(seq_len(nrow(Z)), each = 2), , drop = FALSE]
# Specify prediction matrix Z_pred
Z_pred <- matrix(c(1,0,0,1,1,0,1,0,1), ncol=3, nrow=3, byrow=T)
constrain <- list(shape_coef=factor(c(1, rep(NA, ncol(Z2) - 1))))

ruttergatsonisreg <- fitRutterGatsonisReg(data=RF2,
                                          TP=TP,
                                          FP=FP,
                                          FN=FN,
                                          TN=TN,
                                          study=study,
                                          Z=Z2,
                                          Z_pred=Z_pred,
                                          map=constrain)

ruttergatsonisreg
#>
#> Rutter & Gatsonis Regression Model
#> -----
#>
```

```

#> Number of studies : 47
#> Optimizer          : Converged
#> Hessian            : Positive definite
#> Max |grad|         : 0.0005885859
#> -2 log likelihood : 753.722 ( df = 9 )
#> AIC                : 771.722
#> BIC                : 788.374
#>
#> Use summary() for parameter estimates.
summary(ruttergatsonisreg)
#> $estimates
#>
#>               Estimate Std. Error
#> accuracy_coef  2.45631156 0.32357673
#> accuracy_coef  0.24852208 0.43954066
#> accuracy_coef  0.33141584 0.44260014
#> threshold_coef -0.55005084 0.21334149
#> threshold_coef -0.19625009 0.26096393
#> threshold_coef  0.49585695 0.26223232
#> shape_coef     0.19768191 0.16983950
#> shape_coef     0.00000000 0.00000000
#> shape_coef     0.00000000 0.00000000
#> sigma2_alpha   1.27791200 0.30852400
#> sigma2_theta   0.47684805 0.11344753
#> Lambda_Pred    2.45631156 0.32357673
#> Lambda_Pred    2.70483365 0.32695391
#> Lambda_Pred    2.78772741 0.30577879
#> Theta_Pred     -0.55005084 0.21334149
#> Theta_Pred     -0.74630093 0.20991415
#> Theta_Pred     -0.05419388 0.21213237
#> beta_Pred      0.19768191 0.16983950
#> beta_Pred      0.19768191 0.16983950
#> beta_Pred      0.19768191 0.16983950
#> logitsens      0.82617129 0.28629952
#> logitsens      1.05130416 0.28606887
#> logitsens      1.12639652 0.27689490
#> sens           0.69554476 0.06062743
#> sens           0.74102525 0.05489857
#> sens           0.75517328 0.05119416
#>
#> $sensspec
#>
#>      spec conflevel logitsens Std_Error  CI_Lower CI_Upper      Sens
#> 1 0.8461538      0.95 0.8261713 0.2862995 0.2650345 1.387308 0.6955448

```

```
#> 2 0.8461538      0.95 1.0513042 0.2860689 0.4906195 1.611989 0.7410253
#> 3 0.8461538      0.95 1.1263965 0.2768949 0.5836925 1.669101 0.7551733
#>   SensCI_Lower SensCI_Upper
#> 1      0.5658735      0.8001621
#> 2      0.6202524      0.8336873
#> 3      0.6419166      0.8414559
```

How do I compare models?

In the previous section we fitted the RF data set using the Rutter and Gatsonis subgroup model while keeping the shape parameter equal across all subgroups `constrain="shape"`. Now let's fit the full model allowing for different shape parameters across subgroups and check whether the data lend support to this approach. In addition, let's also consider a model that constrains both `shape` and `accuracy`.

```
ruttermatsonissubfull <- fitRutterGatsonisSubgroup(data=RF2,
                                                    TP=TP,
                                                    FP=FP,
                                                    FN=FN,
                                                    TN=TN,
                                                    study=study,
                                                    subgroup=method,
                                                    constrain=NULL)

ruttermatsonissubfull
#>
#> Rutter & Gatsonis Subgroup Model
#> -----
#>
#> Number of studies      : 47
#> Number of subgroups    : 3
#> Optimizer              : Converged
#> Hessian                : Positive definite
#> Max |grad|             : 0.000337596
#> -2 log likelihood      : 753.553 ( df = 11 )
#> AIC                   : 775.553
#> BIC                   : 795.905
#>
#> Use summary() for parameter estimates.
summary(ruttermatsonissubfull)
#> $estimates
#>                                     Estimate Std. Error
```



```

#> Lambda_LA          2.4243669 0.33049225
#> xi_ELISA           0.2974583 0.51315392
#> xi_Nephelometry    0.3716007 0.45086693
#> Theta_LA          -0.5029487 0.24451662
#> gamma_ELISA        -0.2578242 0.37024676
#> gamma_Nephelometry 0.3970656 0.35944299
#> beta_LA            0.2777491 0.26642343
#> delta_ELISA        -0.1028887 0.42661221
#> delta_Nephelometry -0.1603831 0.39753788
#> sigma2_alpha        1.2730677 0.30798947
#> sigma2_theta        0.4762396 0.11343569
#> Lambda_LA          2.4243669 0.33049225
#> Lambda_ELISA       2.7218252 0.39299597
#> Lambda_Nephelometry 2.7959676 0.30708175
#> Theta_LA          -0.5029487 0.24451662
#> Theta_ELISA        -0.7607729 0.27817185
#> Theta_Nephelometry -0.1058831 0.26322843
#> beta_LA            0.2777491 0.26642343
#> beta_ELISA          0.1748604 0.33321541
#> beta_Nephelometry  0.1173661 0.29500322
#> logitsens          0.8186925 0.27519154
#> logitsens          1.0626991 0.32405869
#> logitsens          1.1206495 0.28834985
#> sens                0.6939587 0.05844518
#> sens                0.7432060 0.06184687
#> sens                0.7541092 0.05346829
#>
#> $sensspec
#>      subgroup      spec conflevel logitsens Std_Error  CI_Lower CI_Upper
#> 1          LA 0.8461538      0.95 0.8186925 0.2751915 0.2793270 1.358058
#> 2          ELISA 0.8461538      0.95 1.0626991 0.3240587 0.4275558 1.697843
#> 3 Nephelometry 0.8461538      0.95 1.1206495 0.2883499 0.5554942 1.685805
#>      Sens SensCI_Lower SensCI_Upper
#> 1 0.6939587 0.5693812 0.7954439
#> 2 0.7432060 0.6052899 0.8452527
#> 3 0.7541092 0.6354094 0.8436717
#>
#> $Reitsma_recovered
#>      mu_A.sens mu_B.spec sigma2_A.sens sigma2_B.spec  sigma_AB
#> LA          0.6172734 1.970652 0.6018282 1.0488717 -0.1579727
#> ELISA        0.5498977 2.315536 0.6670472 0.9463208 -0.1579727
#> Nephelometry 1.2184583 1.594759 0.7065226 0.8934472 -0.1579727

```

```

#>
#> $subgroups
#> [1] "LA" "ELISA" "Nephelometry"
ruttergatsonissubacc <- fitRutterGatsonisSubgroup(data=RF2,
                                                TP=TP,
                                                FP=FP,
                                                FN=FN,
                                                TN=TN,
                                                study=study,
                                                subgroup=method,
                                                constrain=c("shape","accuracy"))

ruttergatsonissubacc
#>
#> Rutter & Gatsonis Subgroup Model
#> -----
#>
#> Number of studies : 47
#> Number of subgroups : 3
#> Optimizer : Converged
#> Hessian : Positive definite
#> Max |grad| : 0.00030144
#> -2 log likelihood : 754.332 ( df = 7 )
#> AIC : 768.332
#> BIC : 781.283
#>
#> Use summary() for parameter estimates.
summary(ruttergatsonissubacc)
#> $estimates
#>
#> Estimate Std. Error
#> Lambda_LA 2.65809857 0.19181608
#> xi_ELISA 0.00000000 0.00000000
#> xi_Nephelometry 0.00000000 0.00000000
#> Theta_LA -0.55745914 0.21293277
#> gamma_ELISA -0.19385254 0.26093654
#> gamma_Nephelometry 0.49702816 0.26213094
#> beta_LA 0.18949665 0.16472998
#> delta_ELISA 0.00000000 0.00000000
#> delta_Nephelometry 0.00000000 0.00000000
#> sigma2_alpha 1.29064650 0.31233907
#> sigma2_theta 0.47671277 0.11344490
#> Lambda_LA 2.65809857 0.19181608
#> Lambda_ELISA 2.65809857 0.19181608

```

```

#> Lambda_Nephelometry  2.65809857 0.19181608
#> Theta_LA            -0.55745914 0.21293277
#> Theta_ELISA         -0.75131168 0.20809954
#> Theta_Nephelometry -0.06043098 0.20958332
#> beta_LA             0.18949665 0.16472998
#> beta_ELISA          0.18949665 0.16472998
#> beta_Nephelometry   0.18949665 0.16472998
#> logitsens          1.00734453 0.16526474
#> logitsens          1.00734453 0.16526474
#> logitsens          1.00734453 0.16526474
#> sens               0.73250015 0.03238258
#> sens               0.73250015 0.03238258
#> sens               0.73250015 0.03238258
#>
#> $sensspec
#>      subgroup      spec conflevel logitsens Std_Error  CI_Lower CI_Upper
#> 1          LA 0.8461538      0.95  1.007345 0.1652647 0.6834316 1.331257
#> 2          ELISA 0.8461538      0.95  1.007345 0.1652647 0.6834316 1.331257
#> 3 Nephelometry 0.8461538      0.95  1.007345 0.1652647 0.6834316 1.331257
#>      Sens SensCI_Lower SensCI_Upper
#> 1 0.7325001  0.6645042  0.7910486
#> 2 0.7325001  0.6645042  0.7910486
#> 3 0.7325001  0.6645042  0.7910486
#>
#> $Reitsma_recovered
#>      mu_A.sens mu_B.spec sigma2_A.sens sigma2_B.spec  sigma_AB
#> LA          0.7018398 2.073994      0.6613828      0.9661567 -0.1540511
#> ELISA        0.5255112 2.287112      0.6613828      0.9661567 -0.1540511
#> Nephelometry 1.1539375 1.527570      0.6613828      0.9661567 -0.1540511
#>
#> $subgroups
#> [1] "LA"          "ELISA"         "Nephelometry"

```

How do I get the log likelihood, the AIC, and the BIC of a model?

```

logLik(ruttergatsonissubfull)
#> 'log Lik.' -376.7765 (df=11)
AIC(ruttergatsonissubfull)
#> [1] 775.5531
BIC(ruttergatsonissubfull)

```

```
#> [1] 795.9047

logLik(ruttergatsonissub)
#> 'log Lik.' -376.8612 (df=9)
AIC(ruttergatsonissub)
#> [1] 771.7223
BIC(ruttergatsonissub)
#> [1] 788.3736

logLik(ruttergatsonissubacc)
#> 'log Lik.' -377.1659 (df=7)
AIC(ruttergatsonissubacc)
#> [1] 768.3318
BIC(ruttergatsonissubacc)
#> [1] 781.2829
```

How do I perform a likelihood ratio test of the two models?

The tests below formally investigate whether the HSROC curve shapes and/or accuracy parameters are equal in all subgroups.

```
anova(ruttergatsonissub,
      ruttergatsonissubfull)
#>           Df  logLik Df.diff  Chisq Pr(>Chisq)
#> Model 1   9 -376.86
#> Model 2  11 -376.78      2 0.1692      0.9189
anova(ruttergatsonissubacc,
      ruttergatsonissub)
#>           Df  logLik Df.diff  Chisq Pr(>Chisq)
#> Model 1   7 -377.17
#> Model 2   9 -376.86      2 0.6095      0.7373
anova(ruttergatsonissubacc,
      ruttergatsonissubfull)
#>           Df  logLik Df.diff  Chisq Pr(>Chisq)
#> Model 1   7 -377.17
#> Model 2  11 -376.78      4 0.7787      0.9413
```

How do I replicate results from the Cochrane Handbook with respect to the schuetz data set?

What are the results for the full data set?

```
data(schuetz)
head(schuetz)
#>   test      study TP FP FN TN indirect
#> 1  CT Achenbach 2005 25  4  0 19         1
#> 2  CT Alkadhi 2008 57 12  2 79         1
#> 3  CT Andreini 2007 17  0  0 44         1
#> 4  CT Bayrak 2008 64  4  0 32         1
#> 5 MRI Bedaux 2002  7  1  0  1         1
#> 6 MRI Bogaert 2003 12  3  3  1         1
schuetz$test <- factor(schuetz$test,levels=c("MRI","CT"))
schuetzreitsma <- fitReitsmaSubgroup(data=schuetz,
                                     TP=TP,FP=FP,FN=FN,TN=TN,
                                     study=study,
                                     subgroup=test,
                                     variances="common")

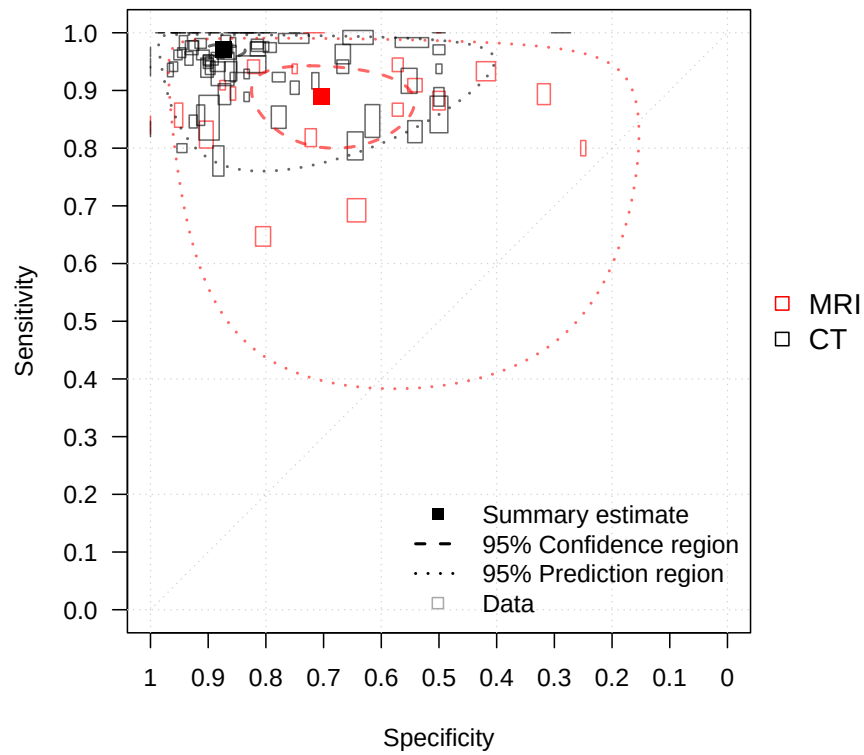
schuetzreitsma
#>
#> Reitsma Subgroup Model
#> -----
#>
#> Number of studies      : 108
#> Number of subgroups   : 2
#> Optimizer              : Converged
#> Hessian                : Positive definite
#> Max |grad|             : 0.0006484896
#> -2 log likelihood      : 951.843 ( df = 7 )
#> AIC                   : 965.843
#> BIC                   : 984.618
#>
#> Use summary() for parameter estimates.
summary(schuetzreitsma)
#> $estimates
#>
#>      Estimate Std_Error
#> mu_A.MRI    2.0934567 0.2639212
#> mu_B.MRI    0.8604695 0.2572264
#> mu_A.CT     3.4827280 0.1569859
```

```

#> mu_B.CT      1.9290688 0.1206664
#> sigma2_A.sens 0.8500392 0.2195458
#> sigma2_B.spec 0.8526826 0.1683378
#> sigma_AB      0.1857463 0.1342067
#> nu_A.CT       1.3893006 0.3015220
#> nu_B.CT       1.0686026 0.2834029
#>
#> $sensspec
#>      type Estimate conflevel CI_Lower CI_Upper
#> mu_A.MRI sens 0.8902656      0.95 0.8286629 0.9315491
#> mu_B.MRI spec 0.7027587      0.95 0.5881481 0.7965102
#> mu_A.CT  sens 0.9701923      0.95 0.9598842 0.9779126
#> mu_B.CT  spec 0.8731463      0.95 0.8445615 0.8971148
#>
#> $RutterGatsonis_recovered
#>      Lambda      Theta      beta sigma2_alpha sigma2_theta
#> MRI  2.954884 0.6176402 0.001552424      2.074212      0.3328068
#> CT   5.413004 0.7789302 0.001552424      2.074212      0.3328068
#>
#> $subgroups
#> [1] "MRI" "CT"
plot(schuetzreitsma,
     nudge_legend=-0.2,
     size="se",
     col=c("red","black"))

```

Diagnostic Test Accuracy Meta-Analysis



```
schuetzreitsma1 <- fitReitsmaSubgroup(data=schuetz,
                                     TP=TP,FP=FP,FN=FN,TN=TN,
                                     study=study,
                                     subgroup=test,
                                     variances="unequal")

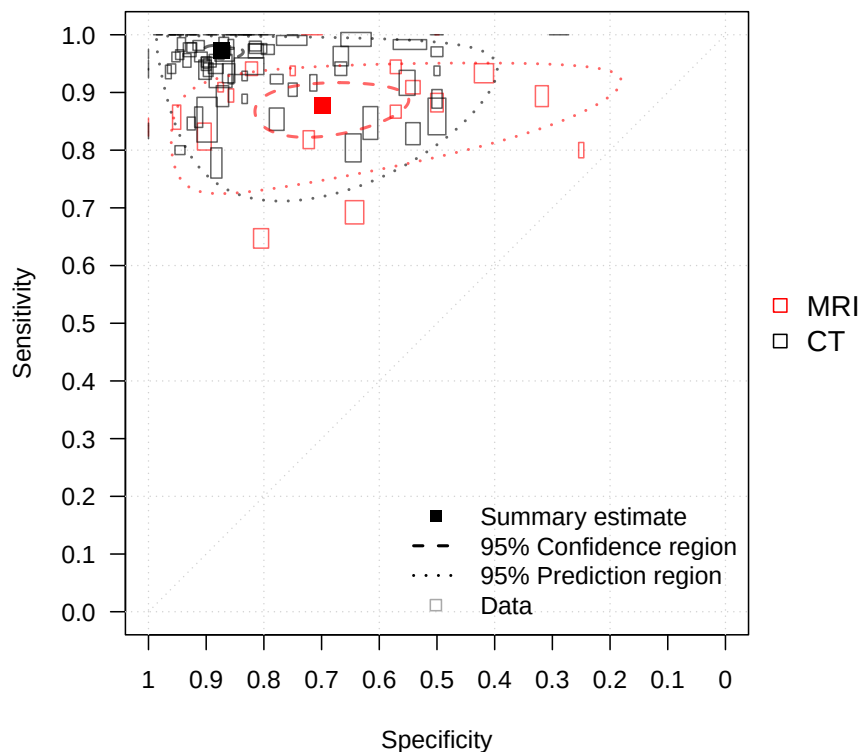
schuetzreitsma1
#>
#> Reitsma Subgroup Model
#> -----
#>
#> Number of studies      : 108
#> Number of subgroups   : 2
#> Optimizer              : Converged
#> Hessian                : Positive definite
#> Max |grad|            : 0.0002796148
#> -2 log likelihood     : 943.305 ( df = 10 )
```

```

#> AIC : 963.305
#> BIC : 990.127
#>
#> Use summary() for parameter estimates.
summary(schuetzreitsma1)
#> $estimates
#> Estimate Std_Error
#> mu_A.MRI 1.9670363 0.1624331
#> mu_B.MRI 0.8410193 0.2406372
#> sigma2_A.MRI 0.1126405 0.1600126
#> sigma2_B.MRI 0.7121292 0.3424753
#> sigma_AB.MRI -0.1462045 0.1370573
#> mu_A.CT 3.5594295 0.1739508
#> mu_B.CT 1.9316450 0.1224654
#> sigma2_A.CT 1.1069322 0.3099921
#> sigma2_B.CT 0.8800412 0.1903877
#> sigma_AB.CT 0.3048890 0.1746537
#> nu_A.CT 1.5923912 0.2379991
#> nu_B.CT 1.0906240 0.2700068
#>
#> $sensspec
#> type Estimate conflevel CI_Lower CI_Upper
#> mu_A.MRI sens 0.8772924 0.95 0.8387117 0.9076606
#> mu_B.MRI spec 0.6986799 0.95 0.5913089 0.7879579
#> mu_A.CT sens 0.9723322 0.95 0.9615243 0.9801668
#> mu_B.CT spec 0.8734314 0.95 0.8444367 0.8976767
#>
#> $RutterGatsonis_recovered
#> Lambda Theta beta sigma2_alpha sigma2_theta
#> MRI 3.649477 1.2943550 0.9220289 0.2740348 0.2147132
#> CT 5.406709 0.6577024 -0.1146895 2.5837546 0.3410496
#>
#> $subgroups
#> [1] "MRI" "CT"
plot(schuetzreitsma1,
     nudge_legend=-0.2,
     size="se",
     col=c("red","black"))

```


Diagnostic Test Accuracy Meta-Analysis



```
schuetzreitsma2 <- fitReitsmaSubgroup(data=schuetz,
                                     TP=TP,FP=FP,FN=FN,TN=TN,
                                     study=study,
                                     subgroup=test,
                                     sensspec_constrain="sens")
schuetzreitsma3 <- fitReitsmaSubgroup(data=schuetz,
                                     TP=TP,FP=FP,FN=FN,TN=TN,
                                     study=study,
                                     subgroup=test,
                                     sensspec_constrain="spec")

anova(schuetzreitsma,schuetzreitsma1)
#>           Df logLik Df.diff  Chisq Pr(>Chisq)
#> Model 1   7 -475.92
#> Model 2  10 -471.65      3 8.5381    0.03611 *
#> ---
#> Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

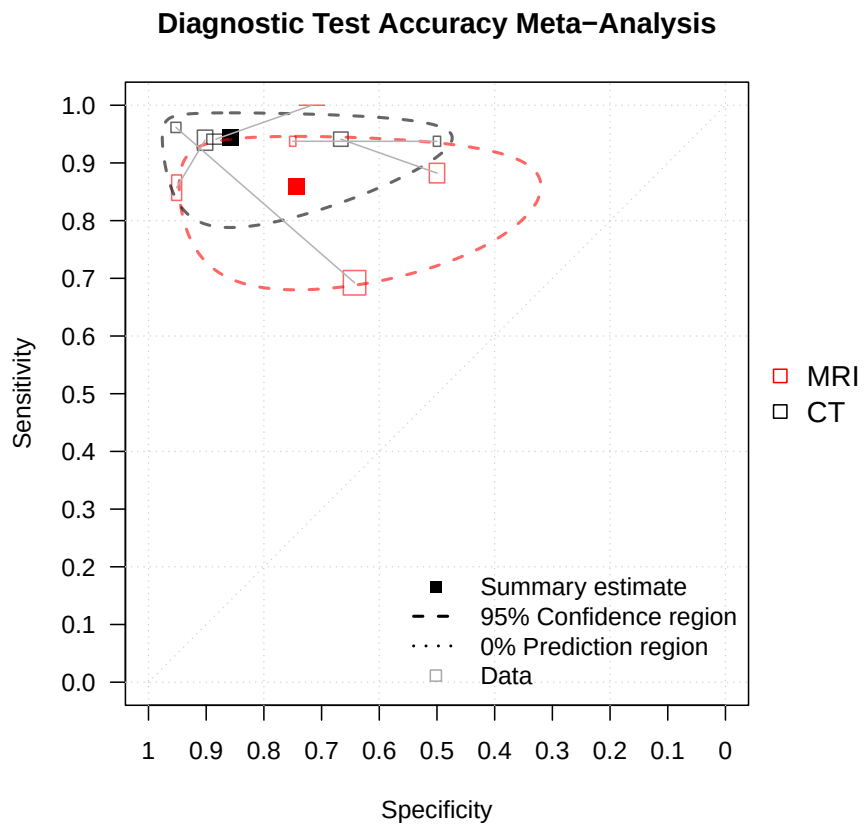
```
anova(schuetzreitsma2,schuetzreitsma)
#>           Df  logLik Df.diff  Chisq Pr(>Chisq)
#> Model 1    6 -485.79
#> Model 2    7 -475.92      1 19.745  8.848e-06 ***
#> ---
#> Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
anova(schuetzreitsma3,schuetzreitsma)
#>           Df  logLik Df.diff  Chisq Pr(>Chisq)
#> Model 1    6 -482.64
#> Model 2    7 -475.92      1 13.432  0.0002474 ***
#> ---
#> Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

What are the results for the direct comparisons?

```
schuetz2 <- subset(schuetz,indirect==0)
schuetzreitsma4 <- fitReitsmaSubgroup(data=schuetz2,
                                     TP=TP,FP=FP,FN=FN,TN=TN,
                                     study=study,
                                     subgroup=test,
                                     constrain="sigma2_A")

schuetzreitsma4
#>
#> Reitsma Subgroup Model
#> -----
#>
#> Number of studies      : 10
#> Number of subgroups   : 2
#> Optimizer              : Converged
#> Hessian                : Positive definite
#> Max |grad|             : 1.86672e-05
#> -2 log likelihood      : 88.734 ( df = 5 )
#> AIC                   : 98.734
#> BIC                   : 100.247
#>
#> Use summary() for parameter estimates.
round(summary(schuetzreitsma4)$estimates,5)
#>
#>           Estimate Std_Error
#> mu_A.MRI      1.80829   0.24124
#> mu_B.MRI      1.06318   0.41604
```

```
#> mu_A.CT      2.81341    0.34319
#> mu_B.CT      1.80140    0.43629
#> sigma2_A.sens 0.00000    0.00000
#> sigma2_B.spec 0.58153    0.43287
#> sigma_AB      0.00000    0.00000
#> nu_A.CT      1.00512    0.41949
#> nu_B.CT      0.73822    0.60560
plot(schuetzreitsma4,predlevel=0.000001,
     nudge_legend=-0.2,
     size="se",scale=0.0025,
     connectstudies = TRUE,
     col=c("red","black"))
```



References

- Reitsma, J. B., et al. (2005). Bivariate analysis of sensitivity and specificity produces informative summary measures in diagnostic reviews. *Journal of Clinical Epidemiology*, 58(10), 982–990.
- Rutter, C. M., & Gatsonis, C. A. (2001). A hierarchical regression approach to meta-analysis of diagnostic test accuracy evaluations. *Statistics in Medicine*, 20(19), 2865–2884.
- Harbord, R. M., Deeks, J. J., Egger, M., Whiting, P., & Sterne, J. A. C. (2007). A unification of models for meta-analysis of diagnostic accuracy studies. *Biostatistics*, 8(2), 239–251.
- Riley, R. D., Ensor, J., Jackson, D., & Burke, D. L. (2018). Deriving percentage study weights in multi-parameter meta-analysis models. *Statistical Methods in Medical Research*, 27(10), 2885–2905.
- Hoyer, A., Hirt, S., Kuss, O. (2018). Meta-analysis of full ROC curves using bivariate time-to-event models for interval-censored data. *Research Synthesis Methods*, 9(1), 62-72.
- Deeks, J. J., Bossuyt, P. M., Leeflang, M. M., & Takwoingi, Y. (editors) (2023). Cochrane Handbook for Systematic Reviews of Diagnostic Test Accuracy. Version 2.0 (updated July 2023). *Cochrane*.